

# MATH 565 Monte Carlo Methods in Finance

Fred J. Hickernell

Test 1

Wednesday, September 25

*Instructions:*

- i. This test consists of THREE questions. Answer all of them.*
- ii. The time allowed is 75 minutes.*
- iii. This test is closed book, but you may use 4 double-sided letter-size sheets of notes.*
- iv. (Programmable) calculators are allowed, but they must not have stored text. No internet access.*
- v. Show all your work to justify your answers. Answers without adequate justification will not receive credit. Write out pseudo-code for the programs that you run to get your answers.*

1. (30 points)

Suppose that  $Y$  is a random variable with finite mean,  $\mu$ , and finite standard deviation,  $\sigma > 0$ . Let  $Y_1, Y_2, \dots$  be IID instances of  $Y$ , and let

$$\bar{Y}_n = \frac{1}{n}(Y_1 + \dots + Y_n)$$

be the sample mean of the first  $n$  samples.

a) Verify that  $\bar{Y}_n$  is unbiased.

*Answer:*

$$\mathbb{E}(\bar{Y}_n) = \mathbb{E}\left[\frac{1}{n}(Y_1 + \dots + Y_n)\right] = \frac{1}{n} [\mathbb{E}(Y_1) + \dots + \mathbb{E}(Y_n)] = \frac{1}{n} [\mu + \dots + \mu] = \mu$$

b) Next compute the variance of  $\bar{Y}_n$ .

*Answer:*

$$\begin{aligned}\text{var}(\bar{Y}_n) &= \text{var}\left(\frac{1}{n}(Y_1 + \dots + Y_n)\right) = \frac{1}{n^2}[\text{var}(Y_1) + \dots + \text{var}(Y_n)] \\ &= \frac{1}{n^2}(\sigma^2 + \dots + \sigma^2) = \frac{\sigma^2}{n}\end{aligned}$$

c) Now consider the estimator  $\bar{Y}_N$ , where the number of samples,  $N$ , is a *random* variable that is independent of the  $Y_i$  with probability mass function

$$\mathbb{P}(N = n) = p_n, \quad n = 1, 2, \dots,$$

for some known  $p_n$ . The laws of conditional probability say that

$$\mathbb{E}(\bar{Y}_N) = \mathbb{E}[\mathbb{E}(\bar{Y}_N|N)], \quad \text{var}(\bar{Y}_N) = \mathbb{E}[\text{var}(\bar{Y}_N|N)] + \text{var}[\mathbb{E}(\bar{Y}_N|N)].$$

In part a) you already computed  $\mathbb{E}(\bar{Y}_N|N = n) = \mathbb{E}(\bar{Y}_n)$ , and in part b) you already computed  $\text{var}(\bar{Y}_N|N = n) = \text{var}(\bar{Y}_n)$ . Use these results to determine whether  $\bar{Y}_N$  is unbiased and also to determine  $\text{var}(\bar{Y}_N)$  in terms of the  $p_n$  above.

*Answer: Since  $\mathbb{E}(\bar{Y}_N|N = n) = \mu$ , independent of  $n$ , it follows that  $\mathbb{E}(\bar{Y}_N|N) = \mu$ , and so*

$$\mathbb{E}(\bar{Y}_N) = \mathbb{E}[\mathbb{E}(\bar{Y}_N|N)] = \mathbb{E}[\mu] = \mu,$$

*so  $\bar{Y}_N$  is unbiased. Also,  $\text{var}(\bar{Y}_N|N = n) = \sigma^2/n$ , and so  $\text{var}(\bar{Y}_N|N) = \sigma^2/N$ , and*

$$\begin{aligned}\text{var}(\bar{Y}_N) &= \mathbb{E}[\text{var}(\bar{Y}_N|N)] + \text{var}[\mathbb{E}(\bar{Y}_N|N)] \\ &= \mathbb{E}\left[\frac{\sigma^2}{N}\right] + \text{var}(\mu) \\ &= \sum_{n=1}^{\infty} \frac{\sigma^2}{n} p_n + 0 = \sigma^2 \sum_{n=1}^{\infty} \frac{p_n}{n}.\end{aligned}$$

2. (30 points)

Suppose that  $Y$  is a random variable with the probability density function  $\rho$  and cumulative distribution function,  $F$ , given by

$$\rho(y) = \frac{e^{-y}}{1 - e^{-1}}, \quad F(y) = \frac{1 - e^{-y}}{1 - e^{-1}}, \quad 0 \leq y \leq 1.$$

Given the *uniform* pseudorandom numbers

$i$	1	2	3	4
$X_i$	0.27850	0.54688	0.95751	0.96489

construct *one* pseudo-random number  $Y_1$  by *either* the inverse distribution transformation method or acceptance-rejection sampling.

*Answer: For the inverse distribution transformation method we first construct the inverse distribution (or quantile) function:*

$$\begin{aligned}x = F(y) = \frac{1 - e^{-y}}{1 - e^{-1}} &\iff (1 - e^{-1})x = 1 - e^{-y} \iff e^{-y} = 1 - (1 - e^{-1})x \\ &\iff y = -\log(1 - [1 - e^{-1}]x) =: F^{-1}(x).\end{aligned}$$

*Taking  $X_1$  above we get*

$$Y_1 = -\log(1 - [1 - e^{-1}]X_1) = -\log(1 - [1 - e^{-1}]0.27850) = 0.19364$$

*(Remember that  $\log$  means natural  $\log$  or  $\ln$  on your calculator.)*

*For acceptance-rejection sampling, we let  $\rho_u$  denote the density function of a  $\mathcal{U}[0, 1]$  random variable, i.e.,  $\rho_u(x) = 1$ , and note that  $(1 - e^{-1})\rho(y) = e^{-y} \leq \rho_u(y)$  for  $0 \leq y \leq 1$ . First we take  $X_1 = 0.27850$ , then we check whether*

$$0.54688 = X_2 \leq (1 - e^{-1})\rho(X_1) = e^{-X_1} = e^{-0.27850} = 0.75692.$$

*Since this is true, we accept 0.27850 as  $Y_1$ .*

3. (40 points)

Consider the situation of  $r = 0$  interest rate. A stock is modeled by a geometric Brownian motion and has an initial price of \$50 and a volatility of 40% (for the whole year). An Asian arithmetic mean call option monitored weekly expires in four weeks. (You may assume 52 weeks per year.) The strike price is \$48.

a) Given the  $\mathcal{N}(0, 1)$  pseudo-random numbers

$$0.71724, -0.35225, -0.37983, -1.76323, \dots$$

construct a single stock path and the Asian arithmetic mean call payoff for this path.

*Answer: The geometric Brownian motion for the stock price is*

$$\begin{aligned} t_j &= \frac{j}{52}, \quad j = 0, \dots, 4 \\ B(0) &= 0 \\ B(t_j) &= B(t_{j-1}) + \sqrt{1/52} X_j, \quad j = 0, \dots, 4, \quad X_1, X_2, \dots \sim \mathcal{N}(0, 1), \\ S(t_j) &= S(0) \exp((r - \sigma^2/2)t_j + \sigma B(t_j)), \quad j = 0, \dots, 4 \\ &= 50 \exp(-0.08t_j + 0.4B(t_j)) \end{aligned}$$

$j$	0	1	2	3	4
$t_j$	0	1/52	1/26	3/52	1/13
$X_j$		0.71724	-0.35225	-0.37983	-1.76323
$B(t_j)$	0.00000	0.09946	0.05061	-0.00206	-0.24657
$S(t_j)$	50.00	51.95	50.87	49.73	45.03

*Thus the payoff is*

$$\max \left( \frac{1}{4} \sum_{j=1}^4 S(t_j) - K, 0 \right) = \max(49.39 - 48, 0) = \$1.39.$$

b) Suppose that the sample mean and sample standard deviation of  $n = 1000$  payoffs are \$2.71 and \$2.93, respectively. Using an approximate central limit theorem (CLT) confidence interval, determine the sample size needed to estimate the option price with an error or no more than \$0.01 and a confidence level of 99%.

*Answer: The half-width of the CLT confidence interval is*

$$\frac{2.58\sigma}{\sqrt{n}} \lesssim \frac{2.58 \times 1.2 \times 2.93}{\sqrt{n}} = \frac{9.10}{\sqrt{n}}.$$

*Thus we choose  $n \approx (9.10/0.01)^2 \approx 830,000$ .*